Mortgage Rates and the Price-to-Rent Ratio Across Space

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Introduction

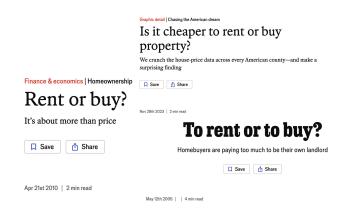
Residential properties can be traded on two markets: **the real estate market and the rental one**.

- ullet Transaction market o transaction price.
- Rental market \rightarrow rent.

What forces drive the **price-to-rent ratios** and their geographic heterogeneity?

- Financial returns for local real estate assets (cap-rate).
- Tenure preferences of households.
- Location preferences through location fundamentals.

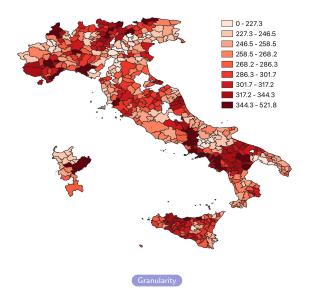
To Buy or to Rent?



Can the price-to-rent ratio convey information on the relative welfare?

Geographic Heterogeneity in Price-to-Rent Ratios

Local price-to-rent ratios Italian local labor markets in 2019.



Financial Returns - Gordon Growth Valuation

The **financial returns** rely on:

- Price = asset price.
- Rent = asset returns.

Thus, under perfectly competitive markets:

$$\mathsf{Price}_{i}^{0} = \sum_{t=0}^{+\infty} \left(\mathsf{Rent}_{i}^{t} * \left(\frac{1}{1 + \mathsf{Rate}_{i}^{t}} \right)^{t} \right)$$

Which is the standard Gordon Growth Valuation formula for assets.

Consistent with structural modeling.

[Vanhapelto, 2022, Amaral, Dohmen, Kohl, and Schularick, 2023, and Greaney, Parkhomenko, and Van Nieuwerburgh, 2025]

Tenure Preferences

Differences between shares of home-owners vs. renters.

Thus different demand for owning and renting \rightarrow **price-to-rent ratio** variation.

Consistent with empirical results.

[Gete and Reher, 2018 and Akgündüz, Dursun-de Neef, Hacihasanoğlu, and Yılmaz, 2023]

This Paper

This paper will:

- Develop a parsimonious model:
 - Housing market with all three forces.
 - ► Augmented Gordon Growth Valuation formula.
- Measure the local responses of prices and rents to a mortgage rate shock:
 - ► The responses of prices and rents vary in magnitude and sign.
 - ► The three forces are necessary to explain the results.
- Run a structural exercise:
 - Recreate the distribution of responses.
 - Show that a mortgage rate hike reduces geographic and tenure inequalities.

Methodology

Common issues when estimating responses to mortgage rate shocks for both prices and rents:

- Difficult to compare properties for sale and for rent.
 - ightarrow Granular Italian dataset which provides local prices and rents for properties of similar quality.
- Endogeneity between prices, rents, and mortgage rates.
 - ightarrow Novel SSIV identification strategy leveraging local demographic differences.

Literature Review

I contribute to three strands of literature:

• Price-to-Rent Ratios:

[Sommer et al. (2013) and Vanhapelto (2022)]

ightarrow General mechanism explaining the local divergence in price-to-rent ratios.

Mortgage Rates and Prices

[Karlman (2022) and Akgündüz et al. (2023)]

 \rightarrow Heterogeneous local responses of prices and rents to mortgage rate shocks.

Quantitative Spatial Models

[Redding and Rossi-Hansberg (2017) and Greaney et al. (2025)]

ightarrow Incorporate the three price-to-rent forces in quantitative spatial models.

Environment

- The model is **static**. There are **I locations**, indexed $i \in \{1, ..., I\}$.
- Households consume a generic consumption good and housing
- Households can either be **renters** or **home-owners**, indexed by $t \in \{owner, renter\}$.
 - \triangleright Renters rent housing by paying a rental price r_i .
 - ► Home-owners buy properties at price *p_i* and decide **how much to consume** and **how much to supply** on the rental market.
 - ► Home-owners enjoy a warm glow bequest motive for owning.
- Households choose location, tenure and consumption.
- Wages and the common consumption prices are given.

Household Problem - Renter

The renter problem is:

$$\max_{\{c_i,H_i^r\}} \ln A_i + \phi_1 \ln c_i + (1-\phi_1) \ln H_i^r + arepsilon_{i,renter}$$
 s.t. $w_i = c_i + r_i H_i^r$ $c_i, H_i^r \geq 0$

- A_i is the value of local amenity in location i.
- ullet c_i is the household consumption choice at price 1
- H_i^r is the housing consumption obtained on the rental market at price r_i .
- $\varepsilon_{i,t} \sim \text{Nested Gumbel}(\sigma_I, \sigma_t)$. Gumbel

Household Problem - Home-Owner

The home-owner problem is:

$$\max_{\{c_i,H_i,H_i^c,H_i^c\}} \, \ln A_i + \ln A_{i,owner} + \phi_1 \ln c_i + \left(1-\phi_1\right) \ln H_i^c + \beta \ln H_i^b + \varepsilon_{i,owner}$$

s.t.
$$w_i + r_i H_i^r = c_i + (1 + \tau) p_i H_i^b$$

 $H_i^b = H_i^r + H_i^c$
 $c_i, H_i^r, H_i^c, H_i^b \ge 0$

- Owners receive an additional local home-ownership amenity $A_{i,owner}$.
- ullet au is a wedge for house prices au Mortgage interest rates.
- β > 0.



Equilibrium

I define the equilibrium as a set $\{p_i, r_i\}_i$ for all $i \in \{1, \dots, I\}$ such that:

- Buyers and renters maximize their utility.
- Builders maximize their profits.
- Transaction housing market clears.
- Rental housing market clears.

Define:

• $\mathsf{Pop}_{i,t} = \lambda_i \, \lambda_{t|i} \, \mathsf{Pop}$: the population which chooses location choice i and tenure choice t.

Equilibrium Price and Rent

Rearranging the market clearing conditions:

Local Rents

$$r_i = (1 - \phi_1) \frac{w_i}{H_i} \left(\mathsf{Pop}_{i,\mathsf{renter}} + \frac{1}{1 + \beta} \mathsf{Pop}_{i,\mathsf{buyer}} \right)$$

Augmented Gordon Growth Valuation Formula

$$(1+ au)p_i = r_i + rac{w_i}{H_i} rac{eta}{1+eta} \mathsf{Pop}_{i,\mathsf{buyer}}$$

What is driving the deviation from the standard Gordon's formula? The **local relative difference in utility**.

Price-to-Rent Ratio

Price-to-Rent Ratio

$$\frac{p_i}{r_i} = \underbrace{\frac{1}{(1+\tau)}}_{1.} + \frac{\beta}{(1-\phi_1)(1+\beta)(1+\tau)} \underbrace{\frac{\mathsf{Pop}_{i,\mathsf{buyer}}}{(\mathsf{Pop}_{i,\mathsf{renter}} + \frac{1}{1+\beta}\mathsf{Pop}_{i,\mathsf{buyer}})}_{2.}}_{2.}$$

The price-to-rent ratio captures the three forces:

- 1. Financial returns: the standard Gordon Growth Valuation formula.
- 2. Location and tenure preferences.

Is the model able to replicate the behavior of prices and rents?

Theoretical predictions

The standard Gordon Growth Valuation formula predicts ($\beta = 0$):

$$\frac{p_i}{r_i} = \frac{1}{1+\tau} \implies \frac{\partial p_i/r_i}{\partial \tau} = -\frac{1}{(1+\tau)^2}$$

Predictions:

- Price-to-rent ratio responses are exclusively negative.
- ullet Net of mortgage rate levels o no price-to-rent ratio responses.

Additionally, without tenure-location equilibrium responses:

- Price responses are **negative**.
- Rent responses are **positive or null** through supply side responses.

OMI Dataset

Empirical analysis of price-to-rent ratios had one issue: **different types** of properties are selected in different markets.

How to compare properties for sale and properties for rent?

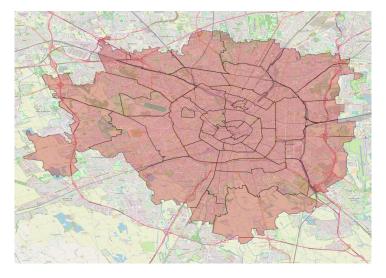
The Italian Tax Agency provides data on prices and rents, based on contract data, aggregated by:

- Sub-municipal homogeneous locations (OMI zones).
- Different categories of residential properties (i.e. 'affordable housing'.)
- Different quality levels.

Thus allowing for a correct comparison between prices and rents.



OMI Zones



Milan OMI zones between 2014 and 2023 overlayed on the city map. Nationally there are >20000 OMI zones.

OLS Estimation - Formula

We aim to estimate the **relationship between mortgage rates and both price and rents**:

$$y_{i,t,q} = \gamma \ln \tau_{R,t} + \mu_m + \mu_q + \mu_{t,LLM} + \varepsilon_{R,t}$$

Where:

- *i* indexes omi zone, *t* is a time index (every semester), *q* indexes the housing quality, *m* indexes municipalities *R* indexes the administrative region.
- $\tau_{R,t}$ is the Italian local mortgage interest rates.
- $\varepsilon_{R,t}$ is the error term, clustered at the R, t level.
- μ_q , μ_m , $\mu_{t,LLM}$ are sets of fixed effect per quality, municipality, and time interacted with local labor markets.



OLS Estimation - Local Heterogeneity Formula

To capture the geographic dimension, I evaluate the OLS regression allowing for different slopes based on local income.

$$y_{i,t,q} = \gamma \ln \tau_{R,t} + \gamma_m W_{m,2012 S1} \ln \tau_{R,t} + \mu_m + \mu_q + \mu_{t,LLM} + \varepsilon_{m,t}$$

• W_m stands for income as measured in 2012.

OLS Estimation Results

Bartik Shift Share Instrument (SSIV)

The parameter of interest is the **response of prices and rents to mortgage rates**.

An SSIV identification strategy requires a variation in **the exposure to a shock**.

Local Demographic Composition

Different age groups have different mortgage pick up rates. \rightarrow **Different local demographic compositions create variation in exposure**.

I further instrument the dependent variable with: a composition of European mortgage interest rates.



SSIV Model - Assumptions

Key Assumptions:

- The shock is quasi-randomly distributed given shares and local unobservables.
 - ▶ The shock is common across each region for every year.
 - The shock is instrumented by a composition of European mortgage interest rates.
- The observations of the effective instrument (shocks) are large in numbers and mutually uncorrelated.
 - Respected across locations and years.

The quasi-random assignment of shocks \implies consistent and unbiased estimation.



SSIV Model - Estimation

I estimate:

$$y_{i,q,t}^{e_m} = \gamma \, \ln \tau_{R,t}^{e_m} + \gamma_m \, W_{m,\, 2012 \, S1} \, \tau_{R,t}^{e_m} + \mu_m + \mu_q^{s_m} + \mu_{t,LLM}^{s_m} + \varepsilon_{m,t}$$

Where:

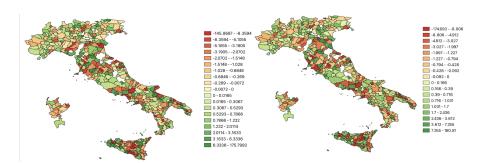
- Superscript e_m states the interaction with the exposure share for municipality m to a mortgage rate shock, s_m with the incomplete sum of age shares.
- $e_m = \sum_a s_{m,a} \kappa_a$, where κ_a is the mortgage pick up rate for age a.
- $au_{R,t}^{e_m}$ is the mortgage interest rate, instrumented by a composition of European mortgage interest rates.

SSIV Analysis - Local Heterogeneity

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	-0.633***	-0.594**	-0.0382
	(0.205)	(0.236)	(0.175)
Interest Rates	0.0526***	0.0458**	0.00682
imes Log of Income	(0.0174)	(0.0197)	(0.0169)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
$Semester \times LLM \; FE$	Yes	Yes	Yes
20 th Percentile	-0.133	-0.159	0.0265
80 th Percentile	-0.113	-0.141	0.0292
N	2'226'042	2'226'042	2'226'042
R-sq	0.9440	0.874	0.937

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the municipality - year level.

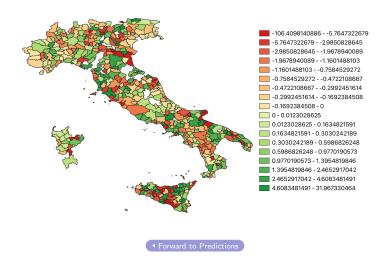
Distributions of Responses - Prices & Rents



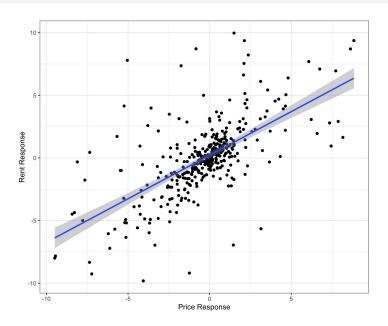
On the left: local price elasticities. On the right: local rent elasticities

Local Labor Markets Distribution of Population Responses

Distribution of Responses - Price-to-Rent Ratios



Price and Rent Local Elasticities Correlation



SSIV Analysis - Local Heterogeneity Additional Results

	Population	Share of Renters
Interest Rates	0.752***	-1.159***
	(0.120)	(0.0453)
Interest Rates	-0.189***	0.00905^*
imes Log of Income	(0.0209)	(0.00482)
Municipality FE	Yes	Yes
$Semester \times LLM \; FE$	Yes	Yes
20 th Percentile	-1.043	-1.073
80 th Percentile	-1.117	-1.070
N	282'242	282'236
R-sq	0.999	0.996

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the municipality - year level.

Pre-Trend Test

Empirical Results

- Prices and rents responses to positive mortgage rate shock are
 - heterogeneous across location.
 - positively correlated within locations.
- Share of renters is less affected in high income locations.

The key model assumptions are **necessary** to account for these results.

- Location/tenure choice.
- Additional ownership utility.

Theoretical Predictions Tested

If $\beta = 0 \rightarrow (1 + \tau)p_i = r_i$. It implies:

$$\frac{p_i}{r_i} = \frac{1}{1+\tau} \implies \frac{\partial p_i/r_i}{\partial \tau} = -\frac{1}{(1+\tau)^2}$$

Which cannot explain:

Positive responses in price-to-rent ratios.

Similarly, absence of location-tenure responses would imply:

- Negative price responses.
- Positive or null rent responses.

Observe **positive correlation** in responses.

Aim of the Structural Estimation

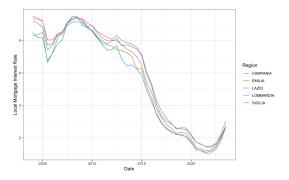
The aim of the estimation is threefold:

- The model generates a similar distribution of prices and rents as the one observed.
- The model generates a set of heterogeneous responses for prices and rents to a mortgage rate hike with respect to local wages as in the empirical estimation.
- Whether there is an heterogeneous local response in welfare for both buyers and renters.

I run a simple counterfactual, isolating the effect of a large **mortgage rate hike** on the economy.

Counterfactual Estimation

The rates of outstanding mortgage loans have increased sharply between 2021 and 2023.



I estimate the price, rents, and welfare effects of the increase in observed mortgage interest rates.

Structural Parameters

I recover the following time-specific parameters for both 2014 and 2021:

- Location amenities A_i .
- Location specific home-ownership amenities A_{owner,i}.
- Location land availability \bar{H}_i .
- Local construction elasticity parameter ρ_i .

I also estimate the following parameters:

- Individual preference parameters ϕ_1 and β .
- Gumbel shock parameters σ_l and σ_t .



Price and Rent Estimation

Does the model generate a realistic price and rent distribution?

Year	Price correlation	Rent correlation
2014	0.217	0.319
2021	0.193	0.361

I report the correlations between

- The model-generated prices and rents.
- The real-world local prices and rents.

Local Responses Distribution w.r.t Local Wages

In the table the direction of the correlation between:

- Model simulated responses and the local observed wages.
- Real-world data estimated responses and the local observed wages

Response	Simulated correlation	Observed correlation
Price	>	>
Rent	>	>
Price-to-Rent		~
(Net of interest rate)		
Population	<	<
Share of Renters	>	>

Local Welfare Responses Distribution

Response	Simulated correlation	
Home-Owner Welfare	<	
Renter Welfare	<	

The responses in local welfare:

- reduce relative spatial inequalities.
- reduce national inequalities between buyers and renters.

Conclusion

The current research finds:

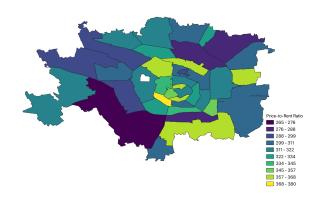
- I incorporate the three forces determining price-to-rent ratios in a spatial model → augmented Gordon Growth Valuation Formula.
- Different locations have different responses in prices and rents w.r.t. mortgage rate shocks → reconcilable only with the three forces.
- The model estimates that the national mortgage rate increase between 2021 and 2023 reduced spatial inequalities and buyers-renters inequalities.

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Thank you for following.

Local Geographic Heterogeneity in Price-to-Rent Ratios

Local price-to-rent ratios across the municipality of Milan in 2019.



Household Problem - Solution

The renter's problem solution:

$$c_{i,\text{Renter}} = \phi_1 w_i, \quad H_{i,\text{Renter}}^r = (1 - \phi_1) \frac{w_i}{r_i}$$

The home-owner's problem solution:

$$c_{i,\mathsf{Owner}} = \frac{\phi_1}{1+\beta} \, w_i, \quad H^c_{i,\mathsf{Owner}} = \frac{1-\phi_1}{1+\beta} \, \frac{w_i}{r_i},$$

$$H_{i,\mathsf{Owner}} = \frac{\beta}{1+\beta} \frac{w_i}{(1+\tau)p_i - r_i},$$

$$H^r_{i,\mathsf{Owner}} = \frac{\beta}{1+\beta} \frac{w_i}{(1+\tau)p_i - r_i} - \frac{1-\phi_1}{1+\beta} \, \frac{w_i}{r_i}$$

Population Distribution

Gumbel idiosyncratic shocks \rightarrow probability of a household to pick tenure choice k given location choice j:

$$\lambda_{k|j} = \frac{\exp(U_{j,k})^{\frac{1}{\sigma_t}}}{\exp(U_{j,k})^{\frac{1}{\sigma_t}} + \exp(U_{j,k'})^{\frac{1}{\sigma_t}}}$$

While the choice of moving to j irrespective of tenure choice is:

$$\lambda_{j} = \frac{\left(\exp(U_{j,k})^{\frac{1}{\sigma_{t}}} + \exp(U_{j,k'})^{\frac{1}{\sigma_{t}}}\right)^{\frac{\sigma_{t}}{\sigma_{l}}}}{\left(\sum_{i=1}^{l} \exp(U_{i,k})^{\frac{1}{\sigma_{t}}} + \exp(U_{i,k'})^{\frac{1}{\sigma_{t}}}\right)^{\frac{\sigma_{t}}{\sigma_{l}}}}$$

◆ Back

Existence of the Rental Market

A location i has a rental market if

$$H_{i,\mathsf{Owner}}^b \geq H_{i,\mathsf{Owner}}^r$$

Rearranging:

$$(1+\beta)r_i \geq (1-\phi_1)(1+\tau)p_i$$

Given that $(1+\tau)p_i \geq r_i$, the existence of a rental market is guaranteed if the relative desire for housing consumption for home-owners (ϕ_2) is lower with respect to the desire for ownership (β) .

Otherwise the rental market collapses and the location has only homeowners with $H^b_{i, \text{Owners}} = H^c_{i, \text{Owners}}$.

Builder Problem

Properties are supplied by a competitive local construction sector.

$$\mathsf{Max}_{\{n_i\}} \ p_i H_i - w_i n_i - p_i^L \bar{H}_i$$

s.t. $H_i = n^{\rho_i} \bar{H}_i^{1-\rho_i}$

- n is the labor input.
- \bar{H}_i is the local land input and p_i^L is its price.
- Under $\rho_i = 0$ the housing supplied is inelastic and equal to \bar{H}_i .



Builder Problem - Solution

The builder's problem solution is:

$$n = \left(\frac{\rho_i \, p_i}{w_i}\right)^{\frac{1}{1-\rho_i}} \bar{H}_i$$

Given that all the available land is traded, the housing supplied is equal to:

$$H_i = \left(\frac{\rho_i \, p_i}{w_i}\right)^{\frac{\rho_i}{1-\rho_i}} \bar{H}_i$$

Data

- House prices and rents: collected by the Italian fiscal agency and aggregated in sub municipal homogeneous territories (OMI zones).
 Provided for houses of similar purposes and quality. Collected between 2004 and 2023.
- Residential properties supply: collected by the Italian fiscal agency and provided in units and number of rooms. Collected between 2013 and 2020.
- Wages: Provided by the Ministry of the Economy and Finance at the municipal level based on income declarations. Collected between 2012 and 2021.
- Population per age: provided at the municipal level by ISTAT. Collected between 2004 and 2023.
- **Share of renters:** provided at the provincial level by ISTAT. Collected between 2004 and 2023.



OLS Estimation

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	0.263	0.271	-0.00805
	(0.228)	(0.235)	(0.145)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
Semester FE	Yes	Yes	Yes
N	2'226'042	2'226'042	2'226'042
R-sq	0.855	0.824	0.620

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the region - year level.

OLS Estimation - Population Weights

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	0.188	0.161	0.0275
	(0.192)	(0.256)	(0.188)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
Semester FE	Yes	Yes	Yes
N	2'226'042	2'226'042	2'226'042
R-sq	0.827	0.803	0.560

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the municipality - year level. Observations are weighted on average local population per OMI zone.

OLS Estimation - Local Heterogeneity

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	-0.119	0.0847	-0.204
	(0.154)	(0.166)	(0.137)
Interest Rates	0.0479***	0.0203	0.0276*
imes Log of Income	(0.0157)	(0.0178)	(0.0152)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
$Semester \times LLM \; FE$	Yes	Yes	Yes
20 th Percentile	0.336	0.277	0.0581
80 th Percentile	0.355	0.285	0.0689
N	2'226'042	2'226'042	2'226'042
R-sq	0.888	0.864	0.754

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the municipality - year level.

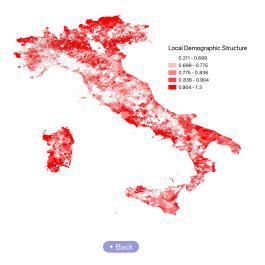
Variation Source - Mortgage Pick Up Rates

Age class	Percentage of borrowers	
Up to 30	64%	
31 - 40	70%	
41 - 50	56%	
51 - 60	36%	
61 - 70	18%	
Over 70	5%	

Distribution of buyers per age class who take a mortgage when buying a residential property. Source: *Quaderni dell'Osservatorio - December 2023*, published by the OMI, based on transaction level data.

Variation Source - Local Demographic Composition

Local municipal population between 18 and 50 divided by the remaining municipal population.



Boryusak et al. (2022) - Assumptions

I consider two assumptions by Borusyak, Hull, and Jaravel, 2022 which allow for a correct estimation of the SSIV model.

Assumption 3: Conditional quasi-random shock assignment

$$\mathbb{E}[g_n|\bar{\varepsilon},q,s]=q_n'\mu,\,\forall n.$$

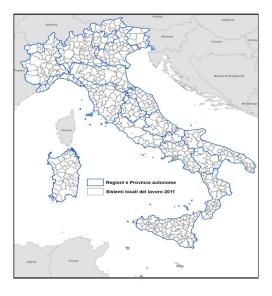
Where g_n is the shock (mortgage interest rates), q_n is a cluster (in our case regions) and μ is a given value.

Assumption 5: Many uncorrelated shock clusters

There exists a partition of shocks into clusters c(n) such that $\mathbb{E}[\sum_c s_c^2] \to 0$ for $s_c = sum_{n:c(n)=c} s_n$ and $Cov[\tilde{g}_n, \tilde{g}_{n'}|\bar{\varepsilon}, q, s] = 0$ for all (n, n') with $c(n) \neq c(n')$.

Where s_n are the exposure shares.

Local Labor Markets in Italy



SSIV - Share definition and Instrument

I define the following:

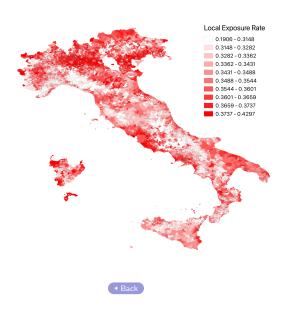
- $s_m = \sum_a s_{a,m}$, sum of the local population shares.
- $e_m = \sum_a s_{a,m} p_a$, linear combination of the local population shares and the mortgage take up rates.

The variable of interest is then: $\tau_{R,t}^{e_m} = \tau_{R,t} e_m$.

Where I index with a the age shares, with m the municipalities, with R the regions and with t the semester. All shares are defined with respect to 2004 S1.

In addition I further instrument the mortgage interest rates with a composition of European national mortgage interest rates.

SSIV - Spatial Distribution of e_m



SSIV - Local Heterogeneity Formula

To test local heterogeneity with respect to income:

$$y_{i,q,t}^{e_m} = \gamma \, \ln \tau_{R,t}^{e_m} + \gamma_m \, W_{m,\, 2012 \, S1} \, \tau_{R,t}^{e_m} + \mu_m + \mu_q^{s_m} + \mu_{t,LLM}^{s_m} + \varepsilon_{m,t}$$

To further assess the local heterogeneity, for each local labor market:

$$y_{i,q,t}^{e_m} = \gamma_{LLM} \, \ln \tau_{R,t}^{e_m} + \mu_m + \mu_q^{s_m} + \, \mu_t^{s_m} + \varepsilon$$



SSIV Analysis - National Level

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	-0.117	-0.146	0.0287
	(0.136)	(0.144)	(0.0543)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
$Semester \times LLM \; FE$	Yes	Yes	Yes
N	2'226'042	2'226'042	2'226'042
R-sq	0.874	0.874	0.937

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors clustered at the region - year level.

◆ Back Population Weighted Regression First Stage Analysis

SSIV Analysis - National Level Population Weighted

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	-0.372**	-0.543***	0.171
	(0.168)	(0.173)	(0.143)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
Semester FE	Yes	Yes	Yes
N	2'226'042	2'226'042	2'226'042
R-sq	0.892	0.843	0.882

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors are clustered at the region - year level. Observation are weighted on local population levels

SSIV Analysis - First Stage

(1)	(2)
0.191***	-11.399***
(0.00917)	(0.130)
	1.391***
	(0.0115)
2'217'900	2'217'900
6'422	80'986
0.997	0.997
	(0.00917) 2'217'900 6'422

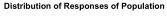
All variables are in log terms. Column (1): national estimations. column (2): estimation accounting for local income differences. '*': significance at the 0.1 level. '**': significance at the 0.01 level. Standard errors are clustered at the OMI zone level.

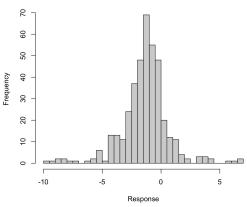
SSIV Analysis - Pre-Trend Local Heterogeneity

	Prices	Rents	Price-to-Rent Ratio
Interest Rates	4.093	3.604	-1.016
	(148899)	(166246)	(39406)
Interest Rates	-0.0581	0.00337	-0.0613
imes Log of Income	(12.962)	(14.471)	(3.44)
Quality FE	Yes	Yes	Yes
Municipality FE	Yes	Yes	Yes
$Semester \times LLM \; FE$	Yes	Yes	Yes
N	51'724	51'724	51'724
R-sq	0.730	0.653	0.739

All variables are in log terms. '*': significance at the 0.1 level. '**': significance at the 0.05 level. '***': significance at the 0.01 level. Standard errors are clustered at the municipality - year level.

Distributions of Responses - Population





SSIV Analysis - Pre-Trend Additional Results

	Population	Share of Renters
Interest Rates	-0.444	-2.532
	(0.980)	(1.498)
Interest Rates	0.0400	0.214
imes Log of Income	(0.0929)	(0.159)
Municipality FE	Yes	Yes
$Year \times LLM \; FE$	Yes	Yes
N	51'724	51'724
R-sq	0.968	0.899

All variables are in log terms. A '*' indicates coefficients significant at the 0.05 level. A '**' indicates coefficients significant at the 0.01 level. A '***' indicates coefficients significant at the 0.001 level. Standard errors are clustered at the municipal - year level.

Dynamic Model with $\beta = 0$

Let us consider the Gordon Growth Formula as stated by Amaral et al., 2023:

$$(1 + \tau_{t=0}) \operatorname{Price}_{i,t=0} = \sum_{t=0}^{+\infty} \frac{\operatorname{Rent}_{i,t}}{(1 + r_{i,t})^t}$$

Where $r_{i,t}$ is the local real interest rate. Assuming rents are expected to grow at constant local rate g_i , that $r_{i,t} = r_i$ for all t, and that $r_i < g_i$:

$$\frac{\mathsf{Price}_i}{\mathsf{Rent}_i} = \frac{1 + g_i}{(r_i - g_i)(1 + \tau)}$$

Thus, even in a simple dynamic model, an increae in mortgage rates that does not affect g_i , cannot explain the observed price-to-rent ratio responses.

Locations Map



I restrict the model to:

- 400 local labor markets.
- The remaining municipalities are aggregated into 20 residual regional locations.



Algorithm

The algorithm is made by the following general steps:

- Set ϕ_1 and ϕ_2 externally.
- Set inner loop to estimate β , σ_l and σ_t .
- Iterate inner loop for A_i , $A_{i,b}$, \bar{H}_i , and ρ_i matching the location choices, tenure choices and the price responses for 2014.
- Recover A_i , $A_{i,b}$, \bar{H}_i , and ρ_i for 2021.
- Run the counterfactual by increasing the interest rates similar to the 2021-2023 mortgage rate hike.



Targeted Moments

- Static shares of:
 - Local population.
 - Local shares of buyers vs. renters
- Responses to a mortgage rate shock (in 2014):
 - Local prices.
- Average magnitude of responses to a mortgage rate shock (in 2014):
 - Local rents.
 - Local population.
 - Local shares of buyers vs. renters

Moments of interest

I am interested in the **correlation of responses** of prices, rents, population and share of buyers with respect to local wages.

